

# Supplemental Notes

Advice: master the iid sum

Thm: If  $X$  and  $Y$  independent and  $Z = X + Y$ :

$$f_z = f_x * f_y$$

"splat" of convolution.

Prf: SIT technique (continuous case)

$$F_z(z) \stackrel{\text{CDF}}{=} P[Z \leq z] \stackrel{\text{Hyp}}{=} P[X + Y \leq z]$$

$$\stackrel{\text{total prob}}{=} \int_{x=-\infty}^{x=\infty} P[X + Y \leq z \mid X=x] \cdot f_x(x) dx$$

$$\stackrel{\text{sub}}{=} \int_{x=-\infty}^{x=\infty} P[x + Y \leq z \mid X=x] \cdot f_x(x) dx.$$

$$\stackrel{\text{ind}}{=} \int_{x=-\infty}^{x=\infty} P[x + Y \leq z] f_x(x) dx$$

$$= \int_{x=-\infty}^{x=\infty} P[Y \leq z - x] \cdot f_x(x) dx.$$

$$\therefore f_z(z) = \frac{d}{dz} F_z(z)$$

$$= \frac{d}{dz} \int_{x=-\infty}^{x=\infty} F_Y(z-x) \cdot f_x(x) dx.$$

$$= \int_{x=-\infty}^{x=\infty} \frac{d}{dz} F_Y(z-x) f_x(x) dx \quad (\text{by Leibniz's Rule})$$

$$= \int_{x=-\infty}^{x=\infty} f_Y(z-x) \cdot f_x(x) dx$$

$$= \left( \underbrace{f_Y * f_x}_{\text{convolve}} \right) (z)$$

then plug-in  $z$

QED

Defn: Moment Generating Function (MBF)

$$M_x(s) = E_x[e^{sX}] \quad \text{for real } s \text{ where } E_x \text{ exists:}$$

(uniform convergence in discrete case) -c < s < c

$s = i\omega$  ( $i = \sqrt{-1}$ ) gives:

Defn: Characteristic function (Lévy, 1920's)

$$\phi_x(\omega) = E_x[e^{i\omega X}]$$

$$|E[e^{i\omega X}]| \leq E[|e^{i\omega X}|] = E[1] = 1$$

Note  $\phi_x(\omega)$  always exists:  $|\phi_x(\omega)| \leq E_x[|e^{i\omega X}|] = \phi_x(0) = 1$   
Jensen Ineq. (triangle)  $\forall \omega \in \mathbb{R}$

Thm: -  $E[X^k] = \frac{d^k}{ds^k} M_x(s) \Big|_{s=0}$  - moment generating

$$- i^k \cdot E_x[X^k] = \frac{d^k}{d\omega^k} \phi_x(\omega) \Big|_{\omega=0}$$

$$\therefore \mu_x = m'_x(0) \quad \sigma_x^2 = m''_x(0) - (m'_x(0))^2$$

Prf:  $k=1$  (continuous case when  $M_x$  exists)

$$\begin{aligned} \frac{d}{ds} M_x(s) \Big|_{s=0} &= \frac{d}{ds} \int_{x=-\infty}^{x=\infty} e^{sx} f_x(x) dx \Big|_{s=0} \\ &= \int_{x=-\infty}^{x=\infty} \frac{d}{ds} e^{sx} \Big|_{s=0} f_x(x) dx \\ &= \int_{x=-\infty}^{x=\infty} x e^{sx} \Big|_{s=0} f_x(x) dx = \int_{x=-\infty}^{x=\infty} x \cdot e^0 f_x(x) dx \\ &= \int_{x=-\infty}^{x=\infty} x \cdot f_x(x) dx = E_x[X] \end{aligned}$$

Repeat  $k$  times for  $E_x[X^k]$ .

Put  $s = i\omega$  for  $\phi_x$

QED

Proposition: If  $X$  and  $Y$  independent and  $Z = X + Y$

$$\left\{ \begin{array}{l} \mathcal{M}_Z = \mathcal{M}_X \mathcal{M}_Y \\ \phi_Z = \phi_X \phi_Y \end{array} \right. \quad \text{n-d: } \mathcal{M}_{\sum_{k=1}^n X_k} = \prod_{k=1}^n \mathcal{M}_{X_k}$$

Prf:  $\mathcal{M}_Z(s) = E_{X,Y} [e^{sX+sY}] = E_{X,Y} [e^{sX} e^{sY}]$

$$\stackrel{\text{ind}}{=} E_{X,Y} [e^{sX}] \cdot E_{X,Y} [e^{sY}] = E_X [e^{sX}] \cdot E_Y [e^{sY}]$$
$$= \mathcal{M}_X(s) \cdot \mathcal{M}_Y(s)$$

QED.

Ex:  $X \sim b(n, p)$  (first) and  $Y \sim P(\lambda)$  (second)

(1)  $\therefore \mathcal{M}_X(s) = E_X [e^{sX}] = \sum_{k=0}^n e^{sk} \binom{n}{k} p^k (1-p)^{n-k}$

$\downarrow$  issue spot

$$= \sum_{k=0}^n \binom{n}{k} (pe^s)^k (1-p)^{n-k}$$

BT

$$= \boxed{(1-p + pe^s)^n}$$

(2)  $\mathcal{M}_Y(s) = E_Y [e^{sY}] = \sum_{k=0}^{\infty} e^{sk} \frac{e^{-\lambda} \lambda^k}{k!}$

$$= e^{-\lambda} \sum_{k=0}^{\infty} \frac{(\lambda e^s)^k}{k!}$$
$$= \boxed{e^{\lambda(e^s - 1)}}$$

★ Thm: (Levy's Continuity Theorem)

(use to prove CLT)  
strategy:  $M_{Z_n} \xrightarrow{c} M_Z \Rightarrow Z_n \xrightarrow{d} Z$

$$X_n \xrightarrow{d} X \quad \text{at points of continuity}$$

$$\text{iff } \phi_{X_n} \xrightarrow{c} \phi_X \quad (\text{and } \phi_X \text{ is continuous})$$

Ex: Poisson Law  $b \xrightarrow{d} p$  if  $\lambda = n \cdot p$

$$\lim_{n \rightarrow \infty} M_X(s) = \lim_{n \rightarrow \infty} (1 - p + pe^s)^n \quad \text{if } X \sim b(n, p)$$

Always remember

$$\sum_{k=0}^{\infty} \frac{x^k}{k!} = e^x$$

$$\lim_{n \rightarrow \infty} \left(1 + \frac{x}{n}\right)^n = e^x$$

$$\stackrel{\lambda=np}{=} \lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n} + \frac{\lambda e^s}{n}\right)^n = \lim_{n \rightarrow \infty} \left(1 + \frac{\lambda(e^s - 1)}{n}\right)^n$$

$$= e^{\lambda(e^s - 1)} = M_Y(s) \quad \text{if } Y \sim P(\lambda)$$

$$\therefore b \xrightarrow{d} p \quad \text{by Levy's Theorem}$$

★ Thm: If  $X \sim N(\mu_x, \sigma_x^2)$

$$= N(0,1) \rightarrow \begin{cases} e^{-s^2/2} \\ e^{-w^2/2} \end{cases}$$

$$1\text{-D} \left\{ \begin{array}{l} M_X(s) = e^{s\mu + s^2\sigma^2/2} \\ \phi_X(w) = e^{i\omega\mu - \omega^2\sigma^2/2} \end{array} \right.$$

memorize

If  $X \sim N(\underbrace{\mu_x}_{n \times 1}, \underbrace{K_{xx}}_{n \times n})$  positive semi-definite

$$n\text{-D} \left\{ \begin{array}{l} M_X(s) = \exp\left(s^T \mu_x + \frac{s^T K_{xx} s}{2}\right) \\ \phi_X(w) = \exp\left(i\omega^T \mu_x - \frac{\omega^T K_{xx} \omega}{2}\right) \end{array} \right.$$

Prf: (1-D case)

$$M_X(s) = E_X[e^{sX}] = \frac{1}{\sqrt{2\pi}\sigma} \int_{x=-\infty}^{x=\infty} e^{sx} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} dx \quad X \sim N(\mu, \sigma^2)$$

put  $z = \frac{x-\mu}{\sigma} \quad \therefore x = \sigma z + \mu \quad dx = \sigma dz$   
 $(x: \pm\infty \rightarrow z: \pm\infty)$

$$= \frac{1}{\sqrt{2\pi}} \int_{z=-\infty}^{z=\infty} e^{s(\sigma z + \mu)} e^{-z^2/2} dz$$

$$= \frac{e^{s\mu}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(z^2 - 2z s\sigma)} dz$$

complete the square

$$= \frac{e^{s\mu}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}[(z-s\sigma)^2 - s^2\sigma^2]} dz$$

since  $(z-s\sigma)^2 - s^2\sigma^2 = z^2 + s^2\sigma^2 - 2zs\sigma - s^2\sigma^2 = z^2 - 2zs\sigma$

$$= e^{s\mu + \frac{s^2\sigma^2}{2}} \left( \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(z-s\sigma)^2} dz \right)$$

put  $u = z - s\sigma \quad \therefore du = dz, \quad z = \pm\infty \rightarrow u = \pm\infty$

$$= e^{s\mu + \frac{s^2\sigma^2}{2}} \left( \frac{1}{\sqrt{2\pi}} \int_{u=-\infty}^{u=\infty} e^{-u^2/2} du \right)$$

= 1 since pdf

$$= e^{s\mu + \frac{s^2\sigma^2}{2}}$$

QED.

Ex: Cauchy  $X \sim C(m, d): f_X(x) = \frac{1}{\pi d(1 + (\frac{x-\mu}{d})^2)}$

$M_X$  does not exist, But  $\phi_X$  does

$$\phi_X(w) = e^{iwm - d|w|}$$

from Cauchy's Residue Theorem:  
 $Z \sim C(0, 1): \phi_Z(w) = e^{-|w|}$

Thm: If independent  $X_1, \dots, X_n$  and  $X_k \sim N(\mu_k, \sigma_k^2)$  then

note: exact

$$\sum_{k=1}^n X_k \sim N\left(\sum_{k=1}^n \mu_k, \sum_{k=1}^n \sigma_k^2\right)$$

↑ similarly distributed

Prf:  $\phi_{\sum_{k=1}^n X_k}(\omega) = E\left[e^{i\omega \sum_{k=1}^n X_k}\right]$

$$= E\left[\prod_{k=1}^n e^{i\omega X_k}\right]$$

ind  $= \prod_{k=1}^n E\left[e^{i(\omega) X_k}\right]$

$$= \prod_{k=1}^n \phi_{X_k}(\omega)$$

$$= \prod_{k=1}^n e^{i\omega \mu_k - \frac{\omega^2}{2} \cdot \sigma_k^2}$$

$$= e^{i\omega \left(\sum_{k=1}^n \mu_k\right) - \frac{\omega^2}{2} \left(\sum_{k=1}^n \sigma_k^2\right)}$$

$$\longleftrightarrow \sum_{k=1}^n X_k \sim N\left(\sum_{k=1}^n \mu_k, \sum_{k=1}^n \sigma_k^2\right)$$

QED.

Q: How does  $\bar{X}_n$  differ if (iid)  $X_k \sim N(\mu, \sigma^2)$  vs.  $X_k \sim C(m, d)$   
sample mean

Ex: If iid  $X_1, \dots, X_n \sim N(\mu, \sigma^2)$ :  $\bar{X}_n \sim N\left(\mu, \frac{\sigma^2}{n}\right)$

Prf:  $\phi_{\bar{X}_n}(\omega) = E\left[e^{i\omega \bar{X}_n}\right] = E\left[e^{i\omega \frac{1}{n} \sum_{k=1}^n X_k}\right]$  (exact, not CLT approximation)

$$= E\left[e^{i \sum_{k=1}^n \frac{\omega}{n} X_k}\right] = E\left[\prod_{k=1}^n e^{i \frac{\omega}{n} X_k}\right]$$

ind  $= \prod_{k=1}^n E\left[e^{i \left(\frac{\omega}{n}\right) X_k}\right] = \prod_{k=1}^n \phi_{X_k}\left(\frac{\omega}{n}\right)$

$X_k \sim N(\mu, \sigma^2)$   
 $= \prod_{k=1}^n e^{i \frac{\omega}{n} \mu_k - \frac{\omega^2}{n^2} \frac{\sigma_k^2}{2}}$

Id.  $= \left(e^{i \left(\frac{\omega}{n}\right) \mu - \frac{\omega^2}{n^2} \left(\frac{\sigma^2}{2}\right)}\right)^n$

$$= e^{i\omega \mu - \frac{\omega^2}{2} \frac{\sigma^2}{n}} \longleftrightarrow \bar{X}_n \sim N\left(\mu, \frac{\sigma^2}{n}\right)$$

QED

★ Ex: If iid  $X_1, \dots, X_n \sim C(m, d)$ :  $\bar{X}_n \sim C(m, d)$   
 $\therefore$  no benefit to averaging!

Prf:  $\phi_{\bar{X}_n}(\omega) = \prod_{k=1}^n E[e^{i(\frac{\omega}{n})X_k}]$  (as in previous proof)

$$= \prod_{k=1}^n \phi_{X_k}\left(\frac{\omega}{n}\right)$$

$$\stackrel{X_k \sim C(m, d)}{=} \prod_{k=1}^n e^{i(\frac{\omega}{n})m - d|\frac{\omega}{n}|}$$

$$= e^{i \sum_{k=1}^n \left( (\frac{\omega}{n})m - d|\frac{\omega}{n}| \right)}$$

$$\stackrel{\text{Id.}}{=} e^{i \left( n \cdot \frac{\omega}{n} \cdot m - n \cdot d \cdot \frac{|\omega|}{n} \right)}$$

$$= e^{i\omega m - d|\omega|} \longleftrightarrow \bar{X}_n \sim C(m, d)$$

QED

$$V[\bar{X}_n] = \frac{\sigma_x^2}{n} \text{ requires } \sigma_x^2 < \infty$$

Note: Francis Edgeworth showed in 1883 that  $\sigma_x^2 < \infty$  is not sufficient, even when  $n=2$ .

★ Inversion (Transform) Theorem

If CDF  $F_x$  is absolutely continuous (so pdf  $f_x$  exists):

$$f_x(\omega) = \frac{1}{2\pi} \int_{-\infty}^{\infty} \phi_x(\omega) \cdot e^{-i\omega x} d\omega$$

# Central Limit Theorem ★★

Facts:  $\sum_{k=0}^{\infty} \frac{x^k}{k!} = e^x$ ,  $\lim_{n \rightarrow \infty} \left(1 + \frac{x}{n}\right)^n = e^x$

Thm: If iid  $X_1, X_2, \dots$  and  $\sigma_x^2 < \infty$  ( $:= |\mu_x| < \infty$ ) then

$$Z_n \xrightarrow{d} Z \sim N(0, 1)$$

where  $Z_n = \text{STD}(\bar{X}_n) = \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}}$   
 $= \text{STD}\left(\sum_{k=1}^n X_k\right) = \frac{\sum_{k=1}^n X_k - n \cdot \mu}{\sqrt{n} \cdot \sigma}$

Prf: Strategy: show  $M_{Z_n} \xrightarrow{c} M_Z(s) = e^{s^2/2}$  and then Levy's theorem.

$$\begin{aligned} M_{Z_n}(s) &= E[e^{s Z_n}] = E\left[e^{s \left(\frac{\sum_{k=1}^n X_k - n \cdot \mu}{\sqrt{n} \cdot \sigma}\right)}\right] \\ &= E\left[e^{\sum_{k=1}^n \frac{s}{\sqrt{n} \cdot \sigma} (X_k - \mu)}\right] \\ &= E\left[\prod_{k=1}^n e^{\frac{s}{\sqrt{n} \cdot \sigma} (X_k - \mu)}\right] \stackrel{\text{ind}}{=} \prod_{k=1}^n E\left[e^{\frac{s}{\sqrt{n} \cdot \sigma} (X_k - \mu)}\right] \\ &= E^n\left[e^{\frac{s(X-\mu)}{\sqrt{n} \cdot \sigma}}\right] \\ &= \left(E\left[1 + \frac{s(X-\mu)}{\sqrt{n} \cdot \sigma} + \frac{s^2(X-\mu)^2}{2n\sigma^2} + \text{H.O.T.}\right]\right)^n \\ &= \left(1 + 0 + \frac{s^2 \sigma^2}{2n\sigma^2} + \text{H.O.T.}'\right)^n \\ &= \left(1 + \frac{s^2/2}{n} + \text{H.O.T.}'\right)^n \xrightarrow{c} e^{s^2/2} \\ &= M_Z(s) \longleftrightarrow Z \sim N(0, 1) \end{aligned}$$

note: need to show H.O.T.'  $\rightarrow 0$  since  $\lim_{n \rightarrow \infty} \left(1 + \frac{c_n}{n}\right)^n = e^c$  if  $\lim_{n \rightarrow \infty} c_n = c$

$\therefore Z_n \xrightarrow{d} Z$  by Levy's continuity theorem, since  $e^{s^2/2}$  is continuous  
QED

## ★ Gamma pdf

$$X \sim \gamma(\alpha, \theta), \quad \therefore f_X(x) = \frac{x^{\alpha-1} e^{-x/\theta}}{\Gamma(\alpha) \theta^\alpha} \quad \text{if } x > 0 \\ (\alpha > 0 \text{ and } \theta > 0)$$

$$\text{for } \Gamma(\alpha) = \int_0^\infty x^{\alpha-1} e^{-x} dx$$

$$\therefore \Gamma(\alpha+1) = \alpha \cdot \Gamma(\alpha) \quad \text{from integration by parts}$$

3-main properties of  $\gamma$  r.vars

Thm: Say  $X \sim \gamma(\alpha, \theta)$

$$\textcircled{1} E[X^k] = \theta^k \frac{\Gamma(\alpha+k)}{\Gamma(\alpha)} \quad \text{if } k > 0$$

moments

(proved week 6)

$$\therefore E[X] = \alpha\theta \quad V[X] = \alpha\theta^2$$

$$\textcircled{2} \text{MGF: } \mathcal{M}_X(s) = \frac{1}{(1-s\theta)^\alpha} \quad \text{if } s < \frac{1}{\theta}$$

Additivity. Say  $X_1, \dots, X_n$  are independent and

$$X_k \sim \gamma(\alpha_k, \theta) \quad \text{Then } X = \sum_{k=1}^n X_k \sim \gamma\left(\sum_{k=1}^n \alpha_k, \theta\right).$$

$$\therefore \sum_{k=1}^r X_k \sim \chi^2(r) \quad \text{if iid } X_k \sim \chi^2(1)$$

$$\text{since } X \sim \chi^2(r) \quad \text{if } X \sim \gamma\left(\frac{r}{2}, 2\right)$$

$$\begin{aligned}
 \textcircled{2} \quad M_x(s) &= E[e^{sX}] = \frac{1}{\Gamma(\alpha)\theta^\alpha} \int_0^\infty e^{sx} x^{\alpha-1} e^{-x/\theta} dx \\
 &= \frac{1}{\Gamma(\alpha)\theta^\alpha} \int_0^\infty x^{\alpha-1} e^{-x(\frac{1}{\theta}-s)} dx \quad \text{if } s < \frac{1}{\theta} \\
 &= \frac{1}{\Gamma(\alpha)\theta^\alpha} \int_0^\infty x^{\alpha-1} e^{-x\frac{(1-s\theta)}{\theta}} dx
 \end{aligned}$$

$$\text{put } u = \frac{x(1-s\theta)}{\theta} \longleftrightarrow x = \frac{\theta u}{1-s\theta} \quad \text{and } dx = \frac{\theta du}{1-s\theta}$$

$$= \frac{1}{\Gamma(\alpha)\theta^\alpha} \int_{u=0}^{u=\infty} \frac{\theta^{\alpha-1} u^{\alpha-1}}{(1-s\theta)^{\alpha-1}} \frac{e^{-u} \theta du}{1-s\theta}$$

$$= \frac{\cancel{\theta^\alpha}}{\cancel{\Gamma(\alpha)\theta^\alpha}} \cdot \frac{\cancel{\Gamma(\alpha)}}{(1-s\theta)^\alpha} = \frac{1}{(1-s\theta)^\alpha} \quad \text{if } s < \frac{1}{\theta}$$

QED (2)

$$\textcircled{3} \quad \text{Say } X_k \sim \mathcal{G}(\alpha_k, \theta) \quad \therefore M_{X_k}(s) = \frac{1}{(1-s\theta)^{\alpha_k}} \quad \text{if } s < \frac{1}{\theta}$$

$$\therefore M_x(s) = E[e^{sX}] = E\left[e^{s\sum_{k=1}^n X_k}\right] = E\left[e^{\sum_{k=1}^n sX_k}\right]$$

$$= E\left[\prod_{k=1}^n e^{sX_k}\right] \stackrel{\text{ind}}{=} \prod_{k=1}^n E[e^{sX_k}] = \prod_{k=1}^n M_{X_k}(s)$$

$$\stackrel{\mathcal{G}}{=} \prod_{k=1}^n (1-s\theta)^{-\alpha_k} = (1-s\theta)^{-\sum_{k=1}^n \alpha_k}$$

$$\longleftrightarrow X \sim \mathcal{G}\left(\sum_{k=1}^n \alpha_k, \theta\right)$$

QED (3)

## CLT and $\chi^2$

$$\sigma\left(\frac{r}{2}, 2\right) \quad \begin{array}{l} \mu_{\chi^2} = r \\ \sigma_{\chi^2}^2 = 2r \end{array}$$

Say  $X \sim \chi^2(r) \quad \therefore X = \sum_{k=1}^r X_k$  with  $\begin{cases} \textcircled{1} \text{ independent } X_1, \dots, X_n \\ \textcircled{2} X_k \sim \chi^2(1) \text{ all } k \end{cases}$

$$\therefore \mu_{X_k} = 1 \quad \text{and} \quad \sigma_{X_k}^2 = 2 \quad (\text{since d.f.} = 1)$$

CLT: iid  $\sum_{k=1}^r X_k \stackrel{d}{\approx} N(r\mu, r\sigma^2)$  for "large"  $r$

$$\therefore X \stackrel{d}{\approx} N(r\mu, r\sigma^2) \quad \text{if } X \sim \chi^2(r) \text{ and large } r \\ = N(r \cdot 1, r \cdot 2)$$

$$\therefore X \stackrel{d}{\approx} N(r, 2r)$$

Similarly:  $\frac{1}{r} X = \frac{1}{r} \sum_{k=1}^r X_k = \bar{X}_r \stackrel{\text{CLT}}{\approx} N\left(1, \frac{2}{r}\right)$